

INTEGRATION OF OIL MARKETS IN INDIA

A. ARUNA KUMARI¹, D. V. SUBBARAO² & K. SUSEELA³

¹Teaching Associate, Department of Agricultural Economics, Agril College, Bapatla, Andhra Pradesh, India

²NABARD Chair Professor, Acharya N G Ranga Agricultural University, Guntur, Andhra Pradesh, India

³Assistant Professor, Department of Agricultural Economics, Agril College, Bapatla, Andhra Pradesh, India

ABSTRACT

The present study was undertaken to assess the market integration of oils in India. Secondary data on monthly wholesale prices were collected for the period of 24 years, i.e., from 1980-81 to 2003-04. The data were collected from Agricultural prices in India. Price series correlation and Augmented Dicky Fuller (ADF) test were used as analytical tools. The results revealed that the selected oil markets for groundnut, rapeseed and mustard, castor and sesame were well integrated at the same order and there is existence of integration of prices between the above oil markets in India. These results have important policy implications in a situation when oil markets are spatially integrated and the government was thinking of reducing or even withdrawing its efforts to influence the price in the market.

KEYWORDS: Market Integration, Oils, Price Series Correlation, Co integration, ADF Test & India

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INTRODUCTION

Although India ranks fourth among the largest producers of oilseeds in the world such as USA, China and Brazil, its productivity is quite low. The low and fluctuating yields are primarily due to a large part of the cultivation being on marginal lands lacking irrigation and with low levels of input usage. Three oilseeds: groundnut, soybean and rapeseed/mustard, together account for over 80 per cent of aggregate cultivated oilseeds output. The edible oils/ oilseeds sector currently faces several challenges. Oilseed cultivation is becoming increasingly unattractive due to low and unstable yields. The marketing system for oilseeds and edible oils has been characterized by government intervention with varying nature and degree of controls and regulations since 1970. A large spread at each level is considered to be a manifestation of market imperfections and poor competitiveness of the oilseed sector. As a result, while growers are deprived of fair returns for their produce, and they face uncertainty in their incomes and production disincentives, the consumers may find it difficult to access the goods at the right time and at reasonable prices (Chand, 2007). In a perfectly competitive market, commodity prices should move in unison in response to the forces of demand and supply. The accuracy and speed at which price change in one market (commodity) gets transmitted to other markets (close substitutes) is taken as an indicator of integration (interdependence) among the markets (commodity groups). The extent of integration gives signals for efficient resource allocation, which is considered essential for ensuring greater market efficiency, price stability and food security. Test of integration also plays a key role in determining the geographical level at which agricultural price policy should be targeted, at least in the short-run to ensure regular availability of food and price stability (Jha et al., 2005; Acharya, 2001). Markets that are not integrated may convey inaccurate price information that might distort marketing decisions and contribute to inefficient producing movements. Hence a study has been undertaken to analyze the integration of oil markets in India.

METHODOLOGY

Integration among the selected oil markets over a period of time i.e., 1980-81 to 2003-04 was examined by employing co integration model. Market integration is defined as a situation in which arbitrage causes prices to move together in different markets. Thus, more specifically, two markets may said to be spatially integrated whenever trade takes place between them, if the price differential for homogenous commodity equals the transfer costs involved in moving that commodity. However, imperfections in the market, particularly those arising from activities of traders are generally taken as important causes of the existence of differential price movements in different markets. Spatial price relationships have been widely used to indicate overall market performance. The usual definition in the literature is that integrated markets are those where prices are determined interdependently. This has generally been assumed to mean that the price changes in one market will be fully transmitted to other markets. Markets that are not integrated may convey inaccurate price information that might distort marketing decisions and contribute to inefficient producing movements. So, it is necessary to compare market prices of oils. For the present study different oil markets, groundnut oil, rapeseed and mustard oil, castor oil and sesame oil were selected. For groundnut oil six markets were selected, namely Hyderabad, Rajkot, Bangalore, Mumbai, Calcutta and Delhi. For rapeseed and mustard oil also six markets were selected, namely Patna, Moga, Rothak, Hapur, Calcutta and Laha. For castor oil, six markets were selected, namely Hyderabad, Bangalore, Mumbai, Chennai, Kanpur and Calcutta. For sesame oil, four markets were selected, namely Hyderabad, Mumbai, Chennai and Varanasi. A bivariate price correlation as well as methodology developed by Engle and Granger (1987) have been used to assess whether oil markets are integrated or not.

Price Series Correlation

Price in one market does vary with the actions of buyers and sellers of other markets. The degree to which price information in one market is related to the process of the price information in other markets can be shown through zero-order correlation matrix of prices in these markets. The approach presumes that with the random price behavior expected of a non-integrated market, the bi variant correlation coefficient of price movements will tend to be zero. Conversely, in a perfectly integrated market, correlation coefficient of price movements is expected to be unity. The simple correlation coefficient for the prices in each pair of selected markets can be estimated by the following formula:

$$r = \frac{\sum (p_{1t} - \bar{p}_1)(p_{2t} - \bar{p}_2)}{\sqrt{\sum (p_{1t} - \bar{p}_1)^2 \sum (p_{2t} - \bar{p}_2)^2}}$$

Where,

r=Simple correlation coefficient

P_1^i =Price of the commodity in the first market at i^{th} point of time

P_2^i =Price of the commodity in the second market at i^{th} point of time

\bar{P}_1 =Mean of prices in the first market

\bar{P}_2 = Mean of prices in the second market

The estimates of correlation coefficients were tested for their significance against zero by using “t” test and against one by Fisher’s “Z” transformation and standard normal distribution test.

Unit Root and Cointegration Test

To examine the price relation between two markets, the following regression model was used.

$$P_{it} = a_0 + a_1 P_{jt} + E_t \quad \text{---} \quad (1)$$

Where P_i and P_j are price series of a specific commodity in two markets i and j and E is the residual term assumed to be distributed identically and independently. Parameter a_0 represents domestic transportation costs, processing costs and sales tax, etc. The test of market integration is straight forward if P_i and P_j are stationary variables. Often; however, economic variables are non-stationary in which case the conventional tests are based towards rejecting the null hypothesis. Thus, before proceeding with further analysis, the stationarity of the variables needs to be checked (Granger and Newbold, 1977)

Stationary series is defined as one whose parameters that describe the series (namely the mean, variance and auto correlation) are independent of time, or rather exhibit constant mean and variance and have an auto correlation that are invariant through time. Thus a series is said to be stationary if its mean and variance are constant over time and the value of the covariance between two time periods depends only on the distance or lag between the two time periods and not on the actual time at which the covariance is computed. Once the non-stationarity of the variables is determined, the next step is to test for the presence of co integrating (long-run equilibrium) relationships between the variables.

To test the univariate price series for stationarity, the Augmented Dickey-Fuller (ADF) test has been applied, which tests the null hypothesis of non-stationarity against an alternative of stationarity.

The standard equation of the ADF test is:

$$\Delta P_t = \alpha + (a-1) P_{t-1} + \sum_{i=1}^p C_i \Delta P_{t-i} + u_t \quad \text{---} \quad (2)$$

The test statistic is simply the “t” statistic, however, under the null hypothesis, it is not distributed as student “t”, but the ratio can be compared with critical values tabulated in Fuller (1976). In estimating equation (2), the H_0 is: P_t is 1(1) which is rejected (in favor of 1(0), if $a-1$ is found to be negative and statistically significant.

The definition of co integration used here is that of Engle and Granger (1987). Before proceeding to test for market integration using the Co integration, analysis, the nature of the integration of the variables needs to be determined. The test involved regressing the first difference of the residual series of residual lagged level and lagged dependent variables is as follows:

$$\Delta E_t = Y_1 + E_{t-1} + \sum_{k=1}^N \alpha_k E_{t-k} + V_t$$

Again the test statistic is the “t” statistic of Y_1 . The critical values are tabulated in Davidson-Mackinnon (1993, Table 20.2, p722) if the estimated “t” value for any of the price series market pair exceeds the critical values it indicates that the price series are co integrated.

RESULTS AND DISCUSSIONS

Price Series Correlation

Groundnut Oil: When calculated using the nominal price series for groundnut oil in six selected markets in India, the bivariate correlation coefficients ranged between 0.87 and 0.95 (Table1). The highest “r” value was observed for the market pair containing Hyderabad and Calcutta and the lowest was observed for the market pair Rajkot and Delhi. The

results indicated that all the price series are efficiently correlated among the selected markets in India.

Table 1: Estimates of Correlation Coefficients for Monthly Wholesale Prices of Groundnut Oil between Pairs of Selected Markets in India (1980-81 through 2003-04)

Markets	Hyderabad	Rajkot	Bangalore	Mumbai	Calcutta	Delhi
Hyderabad	1.0000	0.9445	0.9182	0.9310	0.9501	0.9035
Rajkot		1.0000	0.9159	0.9142	0.9134	0.8748
Bangalore			1.0000	0.9464	0.9134	0.8933
Mumbai				1.0000	0.9346	0.9054
Calcutta					1.0000	0.9120
Delhi						1.0000

Rapeseed and Mustard Oil: The results showed the bi variant correlation coefficients ranged between 0.75 and 0.99 (Table 2). The highest “r” value was observed for the market pair containing Kanpur and Calcutta and the lowest for the market pair containing Haryana and Delhi. The results indicated that all the price series were efficiently correlated among selected markets in India.

Table 2: Estimates of Correlation Coefficients for Monthly Wholesale Prices of Rapeseed and Mustard Oil between Pairs of Selected Markets in India (1980-81 through 2003-04)

Markets	Patna	Moga	Haryana	Kanpur	Calcutta	Delhi
Patna	1.0000	0.8946	0.8721	0.9559	0.9648	0.8686
Moga		1.0000	0.9056	0.9254	0.9249	0.8047
Haryana			1.0000	0.8871	0.8909	0.7548
Kanpur				1.0000	0.9878	0.8789
Calcutta					1.0000	0.8782
Delhi						1.0000

Castor Oil: The results showed the bi variant correlation coefficients ranged between 0.58 and 0.99 (Table 3). The highest “r” value was observed for the market pair containing Madras and Kanpur and the lowest for the market pair containing Bangalore and Calcutta. The results indicated that all the price series were efficiently correlated among selected markets in India.

Table 3: Estimates of Correlation Coefficients for Monthly Wholesale Prices of Castor Oil between Pairs of Selected Markets in India (1980-81 through 2003-04)

Markets	Hyderabad	Bangalore	Mumbai	Madras	Kanpur	Calcutta
Hyderabad	1.0000	0.8171	0.9303	0.9577	0.9510	0.9114
Bangalore		1.0000	0.8444	0.9116	0.9102	0.5828
Mumbai			1.0000	0.9253	0.9231	0.8408
Madras				1.0000	0.9879	0.8184
Kanpur					1.0000	0.8195
Calcutta						1.0000

Sesame Oil: The results showed the bi variant correlation coefficients ranged between 0.70 and 0.96 (Table 4). The highest “r” value was observed for the market pair containing Hyderabad and Chennai and the lowest for the market pair containing Mumbai and Varanasi. The results indicated that all the price series were efficiently correlated among selected markets in India.

Table 4: Estimates of Correlation Coefficients for Monthly Wholesale Prices of Sesame Oil between Pairs of Selected Markets in India (1980-81 through 2003-04)

Markets	Hyderabad	Mumbai	Chennai	Varanasi
Hyderabad	1.0000	0.8718	0.9642	0.7368
Mumbai		1.0000	0.8690	0.7013
Chennai			1.0000	0.7140
Varanasi				1.0000

Augmented Dickey Fuller (ADF) Test

Groundnut Oil: The results of the ADF test were presented in Table 5. The results would show that to the original level, the “t” calculated values were -1.06,-1.19,-1.52,-1.03,-0.25 and -0.92 in Hyderabad, Bangalore, Mumbai, Calcutta and Delhi markets respectively. All these values were less than asymptotic critical value -2.57 under constant and no trend, but at the first difference (n-1 level), these “t” calculated values were -4.54,-7.00,-4.54,-4.49,-4.76 and -4.47 respectively for the same markets, which was greater than the critical value. It indicated that all the price series were stationary and had the same order of integration denoted by 1 (1). The results also showed that all the price series were having the same order of integration, and hence there was no need to go for further tests. If any series is having the same order of integration it may constitute a potential co integration relationship (Gujarati, 1996). Hence there was the existence of integration of prices between the selected groundnut oil markets in India.

Rapeseed and Mustard Oil: The results of the ADF test were presented in Table 6. The results showed that to the original level, the “t” calculated values were -0.81,-2.23,-1.40,-0.95,-0.72 and -1.21 in Patna, Moga, Haryana, Kanpur, Calcutta and Delhi markets respectively. All these values were less than asymptotic critical value -2.57 under constant and no trend, but at the first difference (n-1 level), these “t” calculated values were -5.57,-5.74,-5.29,-4.47,-4.73 and -5.67 respectively for the same markets, which was greater than the critical value. It indicated that all the price series were stationary and had the same order of integration denoted by 1(1). So there was the existence of integration of prices between the selected Rapeseed and Mustard oil markets in India.

Castor Oil: The results of the ADF test were presented in Table 7. The results showed that to the original level, the “t” calculated values were -1.39,-0.31,-1.20,-0.30,-0.58 and -1.16 in Hyderabad, Bangalore, Mumbai, Chennai, Kanpur and Calcutta markets respectively. All these values were less than asymptotic critical value -2.57 under constant and no trend, but at the first difference (n-1 level), these “t” calculated values were -4.41,-5.01,-4.41,-4.47,-3.78 and -4.29 respectively for the same markets, which was greater than the critical value. It indicated that all the price series were stationary and had the same order of integration denoted by 1(1). So there was the existence of integration of prices between the selected castor oil markets in India.

Sesame Oil: The results of the ADF test were presented in Table 8. The results showed that to the original level, the “t” calculated values were -0.65,-0.44, -0.87and -1.74 in Hyderabad, Mumbai, Chennai and Varanasi markets respectively. All these values were less than asymptotic critical value -2.57 under constant and no trend, but at the first difference (n-1 level), these “t” calculated values were -4.03,-4.09, -3.75and -6.13 respectively for the same markets, which was greater than the critical value. It indicated that all the price series were stationary and had the same order of integration denoted by 1 (1). So there was the existence of integration of prices between the selected sesame oil markets in India.

To sum up, under price series correlation, all the selected oil markets, namely groundnut oil, rapeseed and mustard

oil, castor oil and sesame oil in India were well integrated with a high degree of pricing efficiency. The values of correlation coefficients in most cases ranged between 0.8 and 0.9. Under Augmented Dicky Fuller (ADF) test also, the above oil markets in India were well integrated which indicated that all the price series were stationary with the same order of integration denoted by 1 (1).

Table 5: Unit Root Tests for Groundnut Oil Price Series in Selected Groundnut Oil Markets in India-Original First and Second Level (1980-81 through 2003-04)

a) Original Level							
Variable	Lags	Constant, No Trend		Constant, Trend			Inference
		A(1)=0	A(0)= A(1)=0	A(1)=0	A(0)= A(1)= A(2)=0	A(1)= A(2)=0	
Critical value (10 per cent)		-2.57	3.78	-3.13	4.03	5.34	
Hyderabad	5	-1.06	1.51	-3.89	5.72	7.59	Accept H_0
Rajkot	8	-1.19	1.75	-3.59	5.03	6.46	Accept H_0
Bangalore	10	-1.52	2.17	-1.98	2.23	2.33	Accept H_0
Mumbai	11	-1.03	1.93	-3.54	5.16	6.27	Accept H_0
Calcutta	15	-0.25	2.07	-3.28	5.08	5.50	Accept H_0
Delhi	4	-0.92	1.70	-3.50	4.97	6.13	Accept H_0
b) First Difference Level							
Variable	Lags	Constant, No Trend		Constant, Trend			Inference
		A(1)=0	A(0)= A(1)=0	A(1)=0	A(0)= A(1)= A(2)=0	A(1)= A(2)=0	
Critical value (10 per cent)		-2.57	3.78	-3.13	4.03	5.34	
Hyderabad	12	-4.54	10.32	-4.54	6.90	10.33	Reject H_0
Rajkot	10	-7.00	24.54	-6.99	16.30	24.44	Reject H_0
Bangalore	11	-4.54	10.33	-4.59	7.02	10.53	Reject H_0
Mumbai	11	-4.49	10.06	-4.48	6.69	10.04	Reject H_0
Calcutta	15	-4.76	11.33	-4.76	7.59	11.37	Reject H_0
Delhi	16	-4.47	10.02	-4.46	6.65	9.96	Reject H_0
c) Second Difference Level							
Variable	Lags	Constant, No Trend		Constant, Trend			Inference
		A(1)=0	A(0)= A(1)=0	A(1)=0	A(0)= A(1)= A(2)=0	A(1)= A(2)=0	
Critical value (10 per cent)		-2.57	3.78	-3.13	4.03	5.34	
Hyderabad	13	-10.92	59.61	-10.90	39.59	59.39	Reject H_0
Rajkot	6	-11.19	62.55	-11.17	41.56	62.34	Reject H_0
Bangalore	11	-9.50	45.10	-9.48	29.95	44.93	Reject H_0
Mumbai	14	-7.44	27.68	-7.43	18.38	27.57	Reject H_0
Calcutta	13	-6.94	24.12	-6.93	16.04	24.03	Reject H_0
Delhi	5	-10.93	59.76	-10.92	39.73	59.60	Reject H_0

Table 6: Unit Root Tests for Rapeseed and Mustard Oil Price Series in Selected Rapeseed and Mustard Oil Markets in India (1980-81 through 2003-04)

a) Original Level							
Variable	Lags	Constant, No Trend		Constant, Trend			Inference
		A(1)=0	A(0)= A(1)=0	A(1)=0	A(0)= A(1)= A(2)=0	A(1)= A(2)=0	
Critical value (10 per cent)		-2.57	3.78	-3.13	4.03	5.34	
Patna	4	-0.81	1.13	-3.52	4.84	6.43	Accept H_0
Moga	2	-2.23	2.67	-4.18	5.95	8.73	Accept H_0
Haryana	15	-1.40	1.39	-2.26	2.06	2.67	Accept H_0
Kanpur	10	-0.95	1.03	-3.91	5.65	7.87	Accept H_0
Calcutta	10	-0.72	0.90	-3.84	5.61	7.73	Accept H_0
Delhi	2	-1.21	1.27	-4.36	6.85	9.71	Accept H_0
b) First Difference Level							
Variable	Lags	Constant, No Trend		Constant, Trend			Inference
		A(1)=0	A(0)= A(1)=0	A(1)=0	A(0)= A(1)= A(2)=0	A(1)= A(2)=0	
Critical value (10 per cent)		-2.57	3.78	-3.13	4.03	5.34	
Patna	9	-5.57	15.50	-5.62	10.55	15.82	Reject H_0
Moga	10	-5.74	16.46	-5.73	10.93	16.40	Reject H_0
Haryana	16	-5.29	14.01	-5.30	9.38	14.07	Reject H_0
Kanpur	14	-4.47	9.99	-4.50	6.78	10.16	Reject H_0
Calcutta	12	-4.73	11.19	-4.79	7.68	11.51	Reject H_0
Delhi	10	-5.67	16.09	-5.73	10.99	16.47	Reject H_0
c) Second Difference Level							
Variable	Lags	Constant, No Trend		Constant, Trend			Inference
		A(1)=0	A(0)= A(1)=0	A(1)=0	A(0)= A(1)= A(2)=0	A(1)= A(2)=0	
Critical value (10 per cent)		-2.57	3.78	-3.13	4.03	5.34	
Patna	10	-9.40	44.16	-9.38	29.35	44.03	Reject H_0
Moga	14	-7.20	25.90	-7.18	17.20	25.80	Reject H_0
Haryana	13	-7.31	26.75	-7.30	17.76	26.64	Reject H_0
Kanpur	10	-9.08	41.26	-9.07	27.40	41.11	Reject H_0
Calcutta	12	-8.21	33.69	-8.20	22.39	33.58	Reject H_0
Delhi	8	-10.79	58.17	-10.78	38.71	58.06	Reject H_0

Table 7: Unit Root Tests for Castor Oil Price Series in Selected Castor Oil Markets in India (1980-81 through 2003-04)

a) Original Level							
Variable	Lags	Constant, No Trend		Constant, Trend			Inference
		A(1)=0	A(0)= A(1)=0	A(1)=0	A(0)= A(1)= A(2)=0	A(1)= A(2)=0	
Critical value (10 per cent)		-2.57	3.78	-3.13	4.03	5.34	
Hyderabad	14	-1.39	-1.41	-3.02	3.39	4.62	Accept H_0
Bangalore	13	-0.31	2.10	-1.97	3.06	2.51	Accept H_0

Mumbai	13	-1.20	1.78	-2.69	3.20	3.72	Accept H_0
Chennai	15	-0.30	2.00	-2.46	3.39	3.09	Accept H_0
Kanpur	15	-0.58	1.52	-2.84	3.63	4.05	Accept H_0
Calcutta	9	-1.16	1.37	-2.98	-3.46	4.47	Accept H_0
b) First Difference Level							
Variable	Lags	Constant, No Trend		Constant, Trend			Inference
		A(1)=0	A(0)= A(1)=0	A(1)=0	A(0)= A(1)= A(2)=0	A(1)= A(2)=0	
Critical value (10 per cent)		-2.57	3.78	-3.13	4.03	5.34	
Hyderabad	14	-4.41	9.72	-4.42	6.50	9.75	Reject H_0
Bangalore	12	-5.01	12.54	-5.12	8.73	13.09	Reject H_0
Mumbai	15	-4.41	9.69	-4.42	6.51	9.76	Reject H_0
Chennai	16	-4.47	10.00	-4.48	6.68	10.02	Reject H_0
Kanpur	16	-3.78	7.14	-3.77	4.74	7.12	Reject H_0
Calcutta	14	-4.29	9.21	-4.28	6.12	9.18	Reject H_0
c) Second Difference Level							
Variable	Lags	Constant, No Trend		Constant, Trend		Inference	
		A(1)=0	A(0)= A(1)=0	A(1)=0	A(0)= A(1)= A(2)=0	A(1)= A(2)=0	
Critical value (10 per cent)		-2.57	3.78	-3.13	4.03	5.34	
Hyderabad	12	-7.86	30.90	-7.85	20.53	30.79	Reject H_0
Bangalore	15	-7.30	26.68	-7.29	17.74	26.60	Reject H_0
Mumbai	15	-7.20	25.93	-7.19	17.22	25.84	Reject H_0
Chennai	13	-7.42	27.52	-7.41	18.28	27.42	Reject H_0
Kanpur	16	-6.75	22.78	-6.74	15.13	22.70	Reject H_0
Calcutta	14	-6.28	19.74	-6.27	13.11	19.67	Reject H_0

Table 8: Unit Root Tests for Sesame Oil Price Series in Selected Sesame Oil Markets in India (1980-81 through 2003-04)

a) Original Level							
Variable	Lags	Constant, No Trend		Constant, Trend			Inference
		A(1)=0	A(0)= A(1)=0	A(1)=0	A(0)= A(1)= A(2)=0	A(1)= A(2)=0	
Critical value (10 per cent)		-2.57	3.78	-3.13	4.03	5.34	
Hyderabad	16	-0.65	1.16	-4.78	8.35	11.50	Accept H_0
Mumbai	12	-0.44	1.11	-2.75	3.35	3.98	Accept H_0
Chennai	6	-0.87	1.73	-2.86	4.15	4.44	Accept H_0
Varanasi	0	-1.74	1.75	-2.72	2.64	3.71	Accept H_0
b) First Difference Level							
Variable	Lags	Constant, No Trend		Constant, Trend			Inference
		A(1)=0	A(0)= A(1)=0	A(1)=0	A(0)= A(1)= A(2)=0	A(1)= A(2)=0	
Critical value (10 per cent)		-2.57	3.78	-3.13	4.03	5.34	
Hyderabad	12	-4.03	8.12	-4.03	5.43	8.14	Reject H_0
Mumbai	12	-4.09	8.37	-4.13	5.68	8.52	Reject H_0
Chennai	13	-3.75	7.02	-3.77	4.74	7.11	Reject H_0
Varanasi	6	-6.13	18.76	-6.12	12.49	18.73	Reject H_0

c) Second Difference Level							
Variable	Lags	Constant, No Trend		Constant, Trend		Inference	
		A(1)=0	A(0)= A(1)=0	A(1)=0	A(0)= A(1)= A(2)=0	A(1)= A(2)=0	
Critical value (10 per cent)		-2.57	3.78	-3.13	4.03	5.34	
Hyderabad	12	-8.88	39.45	-8.86	26.21	39.30	Reject H_0
Mumbai	12	-8.11	32.87	-8.09	21.84	32.75	Reject H_0
Chennai	13	-7.35	26.99	-7.33	17.93	26.89	Reject H_0
Varanasi	15	-6.67	22.25	-6.66	14.78	22.16	Reject H_0

The results showed that the oil markets in India were well integrated with a high degree of pricing efficiency. The values of correlation coefficients in most cases ranged between 0.8 and 0.9 and it was an indicator of high competition and thus high spatial pricing efficiency. The results are in conformity with the following studies such as Blyn (1973), Thakur (1973), Thakur (1974), Raju and Van Oppen (1982) and Acharya (1988).

Under Augmented Dicky Fuller (ADF) test, the commodities were selected which were similar under price series correlation. The results of this test showed that all oil markets in India were well integrated because the “t” calculated values are greater than asymptotic critical values for all the price series are stationary and having the same order of integration denoted by 1 (1). The results are in conformity with the following studies conducted by Surajit Deb (2004), Pramod Kumar *et al.* (2003), Sunil (2003) and Thomas and Sundaresan (1996).

CONCLUSIONS AND POLICY IMPLICATIONS

All the selected oil markets were well integrated at the same order and there is existence of integration of prices between the selected oils markets in India. These results have important policy implications. In a situation when oil markets are spatially integrated the government may think of reducing or even withdrawing its efforts to influence the price in the market. The finding of the market integration appears to be quite significant for the success of price policy and market liberalization programs undertaken in India.

REFERENCES

1. Acharya S S 1988 *Agricultural Production, marketing and price policy in India*. Mittal Publications. pp:328.
2. Acharya, S S 2001 *Domestic Agricultural Marketing Policies, Incentives and Integration in S. S. Acharya and D. P. Chaudhri (Eds.) (2001), Indian Agricultural Policy at the Crossroads, Rawat Publications, Jaipur.*
3. Blyn G 1973 *Price series correlation as a measure of market integration*. *Indian Journal of Agricultural Economics* 28(2):56-59
4. Chand, Ramesh 2007 *Marketing and Trade Policies in a Globalizing World, 4th Parthasarthy Memorial Lecture delivered during the 21st Conference of Indian Society of Agricultural Marketing, published in Indian Journal of Agricultural Marketing, (Conference Special), 21(3).*
5. Davidson, Russell and Mackinnon J G 1993 *Estimation and inference in econometrics*. Oxford University Press, New York.
6. Engle R F and Granger C W J 1987 *Co integration and error correction: Representation, estimation and testing* *Econometrical* 55:251-276.
7. Fuller W A 1976 *Introduction to stastical time series*. Wiley, New York

8. Granger C W J and New bold P 1977 forecasting economic time series. Academic Press, New York
9. Gujarati and Damoder N 1996 Basic Econometrics. Mc Graw Hill Incorporation, Newyork, and pp: 719.
10. Jha, Raghbendra, K. V. Bhanu Murthy and Anurag Sharma 2005 Market Integration in Wholesale Rice Markets in India ASARC Working Paper 03, Australian National University, Canberra, Australia.
11. Pramod Kumar, Sharma R K and Kumar P 2003 Spatial price integration and pricing efficiency at the farm level: A study of paddy in Haryana. *Indian Journal of Agricultural Economics* 58(2):201-217.
12. Raju V T and Vonoppen M 1982 Marketing efficiency for selected crops in semi arid tropical India. Progress Report No.32, ICRISAT Economics Programme.
13. Sunil A 2003 The process of interactions between industry and agriculture in India: Evidence based on a multivariate co integration and error correction modeling approach. *Indian Journal of Agricultural Economics* 58(4):715-728.
14. Surjit Deb 2004 Terms of trade and investment behavior in Indian agriculture: A co integration analysis. *Indian Journal of Agricultural Economics* 59(2):209-229.
15. Thakur D S 1973 Pricing efficiency of the Indian apple market. *Indian Journal of Agricultural Economics* 28(1):105-114.
16. Thakur D S 1974 Food grain marketing efficiency: A case study of Gujarat. *Indian Journal of Agricultural Economics* 29(4):61-74.
17. Thomas J K and Sundaresan R 1996 Export performance of cardamom in India. *Bihar Journal of Agricultural Marketing* 4(1):29-34.